

Intro

- The term "robust" in filtering and identification refers to the use of inference criteria that are more robust than L^2 norm. They can be considered special case of Huber functions, where the residual is re-weighted, rather than data selected or rejected. More importantly, the inlier/outlier decision is typically instantaneous.

Notation.

- The spatial frame S is attached to Earth and oriented, so gravity $r^T = [0 \ 0 \ 1]^T \|r\|$ is known.
- The body frame b is attached to the IMU.
- The camera frame is c .
- The motion is described in the body frame at time t relative to the spatial frame $g_{sb}(t)$. Since the spatial frame is arbitrary, it is co-located with the body at $t=0$.
- To simplify, $g_{sb}(t) \Rightarrow g$, and omit s_b .

$$\begin{cases} \dot{T} = v & T(0) = 0 \\ \dot{R} = R(\hat{\omega}_{imu} - \hat{\omega}_b) + n_R & R(0) = R_0 \\ \dot{v} = R(\alpha_{imu} - \alpha_b) + r + n_v \\ \dot{\omega}_b = \omega_b \\ \dot{\alpha}_b = \alpha_b \end{cases}$$

gravity $r \in \mathbb{R}^3$ is treated as a known parameter.

ω_{imu} are the gyro measurements

ω_b their biases

α_{imu} the accel measurement

α_b their unknown bias

R_0 the unknown initial orientation of body wrt gravity.

- Initially, we assume points p_i with coordinates $X_i \in \mathbb{R}^3$, $i=1 \dots N$ visible from time $t=t_i$ to current time t .

$\pi: \mathbb{R}^3 \rightarrow \mathbb{R}^2$, $X \rightarrow [\frac{X_1}{X_3}, \frac{X_2}{X_3}]$ is perspective projection.

- A detector and tracker yields $y_i(t)$, $i=1 \dots N$

$$y_i(t) = \pi(g^{-1}(t)p_i) + n_i(t), \quad t \geq 0.$$

where $\pi(g^{-1}(t)p_i)$ is $\frac{R_{1:2}^T(t)(X_i - T(t))}{R_3^T(t)(X_i - T(t))}$ with $g(t) \doteq (R(t), T(t))$.

g_{cb} is body frame to camera frame:

$$y_i(t) = \pi(g_{cb} g^{-1}(t)p_i) + n_i(t) \in \mathbb{R}^2$$

The unknown constant parameters p_i and g_{cb} can then be added to the state with trivial dynamics:

$$\begin{cases} \dot{p}_i = 0, \quad i=1 \dots N(j) \\ \dot{g}_{cb} = 0 \end{cases}$$

State $x = \{T, R, v, \omega_b, \alpha_b, T_{cb}, R_{cb}\}$. where $g = (R, T)$, $g_{cb} = (R_{cb}, T_{cb})$.

and structure parameters p_i are represented in coordinates by

$$X_i = \bar{y}_i(t_i) \exp(p_i)$$

which ensures $z_i = \exp(p_i)$ is possible.

- Define the known input $u = \{\hat{\omega}_{imu}, \alpha_{imu}\} = \{u_1, u_2\}$, the unknown input $v = \{\omega_b, \alpha_b\} = \{v_1, v_2\}$, and the model error $w = \{n_R, n_v\}$. After defining $f(x)$, $c(x)$, matrix D and $h(x, p) = [\dots, \pi(R^T(X_i - T)), \dots]^T$ with $p = \{p_1, \dots, p_N\}$,

$$\begin{cases} \dot{x} = f(x) + c(x)u + Dv + c(x)w \\ \dot{p} = 0 \\ y = h(x, p) + n \end{cases}$$

- To enable a smooth estimate, we augment the state with a delay-line, For a fixed interval dt and $1 \leq n \leq k$, define

$$x_n(t) \doteq g(t - ndt), \quad x^k \doteq \{x_1, \dots, x_k\}$$

that satisfies

$$x^k(t+dt) = Fx^k(t) + Gx(t)$$

where

$$F \doteq \begin{bmatrix} 0 & 0 & 0 \\ I & 0 & 0 \\ 0 & I & 0 \end{bmatrix}, \quad Gx(t) \doteq \begin{bmatrix} g(t) \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$

$$x \doteq \{x_1, x_2, \dots, x_k\} = \{x, x^k\}.$$

A k stack of measurements

$$y_j^k(t) = \{y_j(t), y_j(t-dt), \dots, y_j(t-kdt)\}$$

can be related to the smoother's state $x(t)$ by

$$y_j(t) = h^k(x(t), p_j) + n_j(t)$$

$$h^k(x(t), p_j) \doteq [h(x(t), p_j) \pi(x_1(t)p_j) \dots \pi(x_k(t)p_j)]^T$$

The overall model is

$$\begin{cases} \dot{x} = f(x) + c(x)u + Dv + c(x)w \\ x^k(t+dt) = Fx^k(t) + Gx(t) \\ \dot{p}_j = 0 \\ y_j(t) = h^k(x(t), p_j) + n_j(t), \quad t \geq t_j, j=1 \dots N(t). \end{cases}$$

- As long as gyro and accel bias rates are not zero, convergence of any inference to a unique estimate cannot be guaranteed.

Robust filtering

- Assume all points appear at time $t=0$, and are present at time t , measurements up to time t as $y^t = \{y(0), \dots, y(t)\}$, Inliers p_j , $j \in J$, $J \subset \{1 \dots N\}$, $|J| \ll N$.

- Assume u, v are absent and p_i are known

$$\begin{cases} \dot{x} = f(x) + w \\ y = h(x) + n \end{cases}$$

To determine whether y_i is inlier, consider the event $I \doteq \{i \in J\}$, compute its posterior

given all data: $P[I|y^t]$, and compare with alternative $P[\bar{I}|y^t]$, $\bar{I} \doteq \{i \notin J\}$ using the posterior ratio

$$\begin{aligned} L(i|y^t) &\doteq \frac{P[I|y^t]}{P[\bar{I}|y^t]} \\ &= \frac{P_{in}(y_i^t | y_{-i}^t)}{P_{out}(y_i^t)} \frac{\epsilon}{1-\epsilon} \end{aligned}$$

where $y_{-i} \doteq \{y_j | j \neq i\}$ are all data points but the i -th,

$P_{in}(y_j) \doteq P(y_j | j \in J)$ is inlier density,

$P_{out}(y_j) \doteq P(y_j | j \notin J)$ is outlier density.

$\epsilon \doteq P(i \in J)$ is prior.

- The $P_{in}(y_{J_s}^t)$ for any subset of inlier set $y_{J_s} \doteq \{y_j | j \in J_s \subset J\}$ can be computed recursively at each t :

$$P_{in}(y^t) = \prod_{k=1}^t P(y_k | y^{k-1}).$$

The smoothing state x^t has the property of making future inlier measurements $y_i(t+1)$, $i \in J$ conditionally independent of their past $y_i^t: y_i(t+1) \perp y_i^t | x(t+1)$ and making inliers independent of each other $y_i^t \perp y_j^t | x^t$ $i \neq j \in J$.

Using these independence, factors can be computed via standard filtering.

$$P(y_k | y^{k-1}) = \int P(y_k | x_k) dP(x_k | x_{k-1}) dP(x_{k-1} | y^{k-1})$$

starting from $P(y_1 | \emptyset)$, where the **densifying** $P(x_k | y^k)$ is maintained by KF.

- Conditioned on a hypothesized inlier set J_{-i} , not containing i , the discriminant $L(i|y^t, J_{-i})$ is

$$L(i|y^t, J_{-i}) = \frac{P_{in}(y_i^t | y_{J_{-i}}^t)}{P_{out}(y_i^t)} \frac{\epsilon}{(1-\epsilon)}$$

can be written as

$$L(i|y^t, J_{-i}) = \frac{\int P_{in}(y_i^t | x^t) dP(x^t | y_{J_{-i}}^t)}{P_{out}(y_i^t)} \frac{\epsilon}{1-\epsilon}$$

where $x^t = \{x(0), \dots, x(t)\}$.

The **smoothing density** $P(x^t | y_{J_{-i}}^t)$ is maintained by smoother

- The challenge is that we do not know the inlier set J_{-i} :

$$\begin{aligned} P_{in}(y_i^t | y_{J_{-i}}^t) &= \sum_{J_{-i} \in P_{-i}^N} P(y_i^t | J_{-i} | y_{J_{-i}}^t) \\ &= \sum_{J_{-i} \in P_{-i}^N} P_{in}(y_i^t | y_{J_{-i}}^t) P[J_{-i} | y_{J_{-i}}^t] \end{aligned}$$

where P_{-i}^N is the power set of $\{1 \dots N\}$ not including i .

- To compute the posterior ratio, we have to marginalize J_{-i} , i.e. average over all possible $J_{-i} \in P_{-i}^N$

$$L(i|y^t) = \sum_{J_{-i} \in P_{-i}^N} L(i|y^t, J_{-i}) P[J_{-i} | y^t]$$